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Cover sheet

Title

Idiopathic pulmonary fibrosis is associated with common genetic variants and limited rare variants

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Idiopathic pulmonary fibrosis is associated with common genetic variants and limited rare variants

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Impact: This is the first study to comprehensively assess the impact of rare variants in idiopathic pulmonary fibrosis, using an agnostic analysis strategy. These results have advanced our understanding of IPF genetics by highlighting the etiologic importance of only two well-established rare genetic variants (*TERT* and *RTEL1*), replicating common variants, and defining the heritability of IPF. In aggregate, these findings simplify the genetics of IPF.

This article has an online data supplement.

ABSTRACT

Rationale

Idiopathic pulmonary fibrosis is a rare, irreversible, and progressive disease of the lungs.

Common genetic variants, in addition to non-genetic factors, have been consistently associated with IPF. Rare variants identified by candidate gene, family-based, and exome studies have also been reported to associate with IPF. However, the extent to which rare variants genome-wide may contribute to the risk of IPF remains unknown.

Objectives

We used whole-genome sequencing to investigate the role of rare variants, genome-wide, on IPF risk.

Methods

As part of the Trans-Omics for Precision Medicine Program, we sequenced 2,180 cases of IPF. Association testing focused on the aggregated effect of rare variants (minor allele frequency ≤ 0.01) within genes or regions. We also identified individual variants that are influential within genes and estimated the heritability of IPF based on rare and common variants.

Measurements and Main Results

Rare variants in both *TERT* and *RTEL1* were significantly associated with IPF. A single rare variant in each of the *TERT* and *RTEL1* genes was found to consistently influence the aggregated test statistics. There was no significant evidence of association with other previously reported rare variants. The SNP-heritability of IPF was estimated to be 32% (s.e. 3%).

Conclusions

Rare variants within the *TERT* and *RTEL1* genes and well-established common variants have the largest contribution to IPF risk overall. Efforts in risk profiling or development of therapies for IPF that focus on *TERT*, *RTEL1*, common variants, and environmental risk factors are likely to have the largest impact on this complex disease.

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Key Words: Whole Genome Sequencing, Interstitial Lung Disease, TOPMed, Genetic Association Studies, Telomerase

INTRODUCTION

Idiopathic pulmonary fibrosis (IPF) is a rare, irreversible, and progressive disease of the lungs that affects an estimated 5 million individuals worldwide and is associated with a median survival of 3-5 years.(1-3) IPF is associated with older age, cigarette smoking, and other environmental exposures.(4, 5) In addition, both rare mutations and common genetic variants are reported to contribute to the etiology of IPF, although the heritability of the disease remains unknown. Within 11p15, we discovered a gain-of-function(6) promoter variant in *MUC5B* (rs35705950) that is the dominant genetic risk factor for IPF, present in >50% of affected patients.(7, 8) Genome-wide association studies have identified and validated several other genetic loci with more moderate associations with IPF risk, including genes involved in telomerase maintenance, host defense, and cell-cell adhesion.(7-15) Candidate gene studies, family-based studies, and exome sequencing analyses have also identified rare mutations that associate with IPF.(16-26) However, critical unresolved questions concerning the genetics of IPF remain, including: 1) the extent and types of rare variants genome-wide that contribute to risk, 2) relative contribution of rare vs. common variants to risk, and 3) the genetic heritability of IPF. To address these questions, we have conducted a whole-genome sequencing study of IPF with genome-wide analysis of rare variation.

METHODS

Study Populations and Sequencing

To comprehensively investigate the role of rare variants in the development of IPF, we collected DNA from patients that were diagnosed with IPF according to criteria established by the

American Thoracic Society/European Respiratory Society from institutions across the United States, Europe, and Australia (Table E1). Whole-genome sequencing was performed by the National Institutes of Health (NIH) Trans-Omics for Precision Medicine (TOPMed) Program.(27) Subjects with IPF were compared to out-of-study controls selected from other TOPMed study populations, and identified as unaffected (without evidence of interstitial lung disease). Within the TOPMed program, sequencing of our case and control populations was performed at multiple centers (Table 1). TOPMed sequencing data from the IPF samples were made available in separate “data freezes”, which we used to define our discovery and validation case populations. Samples that were included in TOPMed Freeze 8 were used for the discovery phase of the analysis, in which 1,264 IPF cases were compared to 1,257 unaffected controls selected from the COPDGene study (Table 1). The validation cohort comprised 916 IPF cases and 1,200 unaffected controls that were selected from the FHS and MESA studies, and included in TOPMed Freeze 10. Since the vast majority of our IPF cases were non-Hispanic white, we filtered our case and control samples to those with European ancestry using ancestry informative principal components, in order to minimize population stratification (see Supplemental Methods).

Statistical Analysis

We conducted a preliminary analysis of common genetic variants with a minor allele frequency (MAF) >0.01 , genome-wide, using the combined discovery and validation cohorts. For our primary analyses of rare variants, we used SKAT-O(28) to conduct association testing of the aggregated effect of rare variants within genes or regions, defined as those with a MAF ≤ 0.01 . All analyses were adjusted for sex as a covariate in the models, as well as principal components

of genetic ancestry to control for any residual fine-scale population stratification. We used a Bonferroni correction for the effective number of tests (K_{eff}) in each analysis, which is based on the estimated minimum achievable p-value for each test.(28) Any gene or window-based variant sets with a $p\text{-value} < 0.05/K_{\text{eff}}$ were considered genome-wide significant and included in validation testing (1.76×10^{-5} , 3.22×10^{-6} , 3.24×10^{-8} for the LOF, LOF/missense, and window-based analyses respectively). Our primary analysis strategy included loss-of-function (LOF) variants aggregated within gene-based sets. For this primary analysis, gene sets with a $p\text{-value} < 5 \times 10^{-4}$ were also included in validation testing based on moderate association. Our pre-specified secondary analyses included LOF and missense variants aggregated within gene-based sets, and comprehensive testing of all rare variants aggregated into non-overlapping windows across the genome based on spatial clustering.(29) Variant sets that had been previously reported in the literature and were moderately associated with IPF in our secondary analyses ($p\text{-value} < 5 \times 10^{-5}$ for missense variant analysis or $p\text{-value} < 5 \times 10^{-7}$ for window-based analysis), were also included in validation testing. We used a Bonferroni correction to assess significance in the validation cohort, adjusting for a total of 9 tests ($p\text{-value} < 5.5 \times 10^{-3}$). We used the Rare Variant Influential Filtering Tool (RIFT)(30) to identify individual variants that had a strong influence on the aggregated test statistic for variant sets that were significantly associated with IPF in the validation cohort. For each analysis strategy, we performed a meta-analysis combining statistics from the discovery and validation cohorts. Finally, we used a genome-based restricted maximum likelihood method (GREML)(31) to estimate SNP-heritability in the combined dataset of discovery and validation samples, using all measured variants. Additional details of the

sample selection, variant filtering, and statistical methods can be found in the Supplemental Methods.

RESULTS

In our preliminary analysis of common variants, using the discovery and validation cohorts combined, we observed genome-wide significant ($p\text{-value} < 5 \times 10^{-8}$) associations with loci previously identified in genetic studies of IPF, including variants in *MUC5B*, *TERT*, *TERC*, *DSP*, and others (Table E2). In addition, we identified two novel associations between IPF and variants in the third intron of *MCL1* (OR=0.77; 95% CI 0.71-0.84; $p\text{-value} = 6.41 \times 10^{-9}$) and the first intron of RNA gene ENSG00000260803 (OR=1.72; 95% CI 1.42-2.08; $p\text{-value} = 3.12 \times 10^{-8}$).

The findings from this common variant analysis are consistent with previous genome-wide association studies, and validate the utility of our case and control populations.

In our primary rare variant analysis, which only included loss-of-function (LOF) variants aggregated into gene-based sets, none of the genes met criteria for genome-wide significance, but there were five genes that met our criteria for moderate association in the discovery cohort (Table 2; Figure 1). *ALOX15B* and *RTEL1-TNFRSF6B* (a read-through transcription between *RTEL1* and *TNFRSF6B*) were most strongly associated with IPF ($p\text{-value} = 2.81 \times 10^{-5}$ and $p\text{-value} = 3.49 \times 10^{-5}$, respectively). Rare variants in the *RTEL1* ($p\text{-value} = 1.11 \times 10^{-4}$), *UNC93A* ($p\text{-value} = 3.44 \times 10^{-4}$), and *NFX1* ($p\text{-value} = 4.67 \times 10^{-4}$) genes were also moderately associated with IPF. These five genes were tested in our independent validation cohort of 916 IPF cases and 1,200 unaffected controls of European ancestry using a Bonferroni $p\text{-value}$ threshold for significance that was corrected for a total of nine tests ($p < 5.5 \times 10^{-3}$). In the validation analysis, only the *RTEL1* gene

was statistically significant after adjustment for multiple testing (p -value= 2.53×10^{-3}). None of the other associations from our LOF analysis strategy replicated within the validation cohort; however, *NFX1* was nominally significant (p -value=0.03).

In a prespecified secondary analysis that included missense variants in addition to the LOF variants, aggregated into gene-base sets, *TERT* (p -value= 3.25×10^{-16}) and *RTEL1* (p -value= 7.49×10^{-9}) were both strongly associated with IPF, exceeding the criteria for genome-wide significance in our discovery cohort (Table 2; Figure 2). The third strongest association signal in this analysis was the *SPDL1* gene (p = 2.73×10^{-5}). Since a rare missense mutation within the *SPDL1* gene has been previously reported to be associated with IPF,(32) we included this gene in our validation testing of rare LOF and missense variants along with *TERT* and *RTEL1*. In the validation cohort, the association with *TERT* was replicated with p -value= 9.39×10^{-8} . The associations with *RTEL1* and *SPDL1* did not reach our Bonferroni-corrected significance threshold for the validation cohort, but were nominally associated with IPF (p -value= 2.07×10^{-2} and p -value= 1.63×10^{-2} , respectively).

In our final analysis of all rare variants, spatially aggregated within non-overlapping windows, none reached our genome-wide significance threshold. At a more moderate level of significance, a ~3200bp window (chr11:1284193-1287389) within the *TOLLIP* gene at 11p15 was associated with IPF (p -value= 1.45×10^{-7} ; Table 2). Since previous studies have reported an association between IPF and variants within the *TOLLIP* gene,(10) we tested this window for association within the validation cohort, where the strength of the association was similar (p -value= 2.89×10^{-6}). Given the proximity of this region to the influential gain-of-function polymorphism within the promoter of the *MUC5B* gene, rs35705950, we repeated the test of

association, adjusting for the *MUC5B* variant. After adjustment for the *MUC5B* promoter polymorphism, the window within *TOLLIP* was no longer associated with IPF (p-value=0.79).

The results of a meta-analysis of the discovery and validation cohorts largely reflected the findings from the individual cohorts (Table 3; Figure 3). In the meta-analysis of LOF variants, *RTEL1* had the strongest association signal, with a p-value just below the threshold for genome-wide significance (p-value=4.25x10⁻⁶), followed by *RTEL1-TNFRSF6B* (p-value=9.12x10⁻⁵), *SPSB2* (p-value=1.02x10⁻⁴), and *PARN* (p =1.43x10⁻⁴). In the LOF/missense variant meta-analysis, both *TERT* and *RTEL1* reached genome-wide significance levels. We tested the association with *TERT*, adjusting for the previously identified common IPF risk variant rs4449583 within *TERT*. (13) The aggregate test statistic remained significant (p-value=3.47x10⁻²¹), indicating that rare variants within *TERT* influence IPF risk independent of the effect of this common *TERT* variant. We also tested the association with *RTEL1* after adjusting for the recently identified IPF risk variant rs41308092 (33). The aggregate test statistic for *RTEL1* remained significant (p-value=6.11x10⁻¹¹). The evidence for association with *SPDL1* was just below the level of genome-wide significance for the meta-analysis. In the window-based analysis, the window within *TOLLIP* was significantly associated with IPF, but not independent of the *MUC5B* variant, as demonstrated by the model adjusting for the rs35705950 genotype (p-value=1.0). None of the other p-values for rare variant associations changed substantially after adjustment for previously identified common IPF risk variants that reside on the same chromosome or the *MUC5B* promoter variant (Table 3).

We did not adjust for age in our analyses, since age was missing for >10% of our cases. However, we performed a sensitivity analysis where association testing was repeated in the

discovery cohort with age as a covariate. Interestingly, we found that the p-values for *TERT*, *RTEL1*, and *SPDL1* decreased after adjustment for age. The other p-values increased slightly, to a degree expected based on the reduced sample size (Table E4).

Based on a previous study that found IPF patients without the *MUC5B* risk (T) allele at rs35705950 had a higher burden of rare missense or LOF variants in *TERT* than those without the risk allele,(34) we examined the frequency of rare variants in *TERT* and *RTEL1* among IPF cases within strata defined by carriage of the *MUC5B* risk allele (GG vs. GT/TT). We did not find a significant difference in the burden of rare LOF and missense alleles in *TERT*. However, the burden of rare LOF and missense alleles in *RTEL1* was greater in cases without the risk allele than cases carrying one or more copies of the risk allele (0.003 vs. 0.001, p-value=0.02).

We applied a recently developed statistical method, the Rare Variant Influential Filtering Tool (RIFT),(30) to identify variants within the *RTEL1*, *TERT*, and *SPDL1* variant sets that had a strong influence on the aggregate test statistic. A single variant in the *RTEL1* LOF variant set, rs373740199, was classified as influential in both the LOF and LOF/missense variant sets, and in both cohorts (Figure E1). This variant is within the 30th exon of *RTEL1*, and was previously identified in an exome sequencing study.(26) The minor allele was present at a frequency of 0.17% among IPF cases and absent among controls. A previously reported IPF risk variant in *TERT*,(24) rs199422297, was influential in the *TERT* LOF/missense variant set across cohorts (Figure E2). This is a stop-gain variant within the 5th exon of *TERT*, and the minor allele was present at a frequency of 0.25% among IPF cases, and absent among controls. These influential *TERT* and *RTEL1* variants are reported in dbSNP to be rare among Europeans (MAF<0.01%) and absent in other populations. A single variant, rs116483731, first identified by exome-wide

association,(32) was also classified as influential in the *SPDL1* LOF/missense variant set across cohorts (Figure E3). This variant is in the second exon of *SPDL1*, and the minor allele frequency was 2.2% among cases, and 0.8% among controls (OR=2.86; 95% CI 1.96-4.17). In dbSNP, the MAF is reported as 0.7% among Europeans, 0.07% among Africans, and is absent among other populations. While other variants may have contributed to the aggregate association test statistics, and to the overall risk of IPF, our analyses suggest that a single rare variant is largely responsible for observed associations in each of the *RTEL1*, *TERT*, and *SPDL1* genes.

We compared the minor allele counts of the three identified rare, influential variants in *TERT*, *RTEL1*, and *SPDL1* in cases with and without a family history of disease. Among IPF cases with non-missing family history data, there were 1,065 sporadic cases and 837 cases with a family history of disease included in our analyses. There was no difference in the proportion of familial and sporadic IPF cases carrying the identified influential minor alleles in *RTEL1* or *SPDL1*. While the minor allele of the influential rare variant in *TERT*, rs199422297, was observed among both sporadic and familial cases, the cases with a family history of disease were more likely to carry the *TERT* minor allele (2/1,065 sporadic vs. 8/837 familial; p-value=0.03).

Finally, we used a genome-based restricted maximum likelihood method, GREML-LMDS,(31) to estimate the heritability of IPF. Using whole-genome sequence data from the combined discovery and validation cohorts, we estimated the SNV-heritability of IPF to be 32% (s.e. 3%).

DISCUSSION

Our findings indicate that rare variants in *RTEL1*, *TERT*, and likely *SPDL1* contribute to the risk of IPF. While these genes have been reported by others to contain rare variants associated with IPF, (13, 17, 20, 23-26, 32) we have found that a single rare variant in each of the implicated genes (*RTEL1*, *TERT*, and *SPDL1*) could be largely responsible for the observed associations. Our whole-genome sequence analysis also suggests that rare variants identified in more focused studies of familial pulmonary fibrosis, including *TERC*, (17) the surfactant protein genes, (16, 18, 21, 22, 35, 36) *TINF2*, (37) and *ABCA3* (38) do not appear to substantially contribute to the overall risk of IPF, at least in a sample of this size (Table E3). In order to further assess any potential effect of individual, previously reported rare variants within our combined dataset, we tabulated the number of minor alleles observed among cases and controls (Table E5). Although additional exceedingly rare variants may prove to be risk factors in unique families or relevant to specific IPF subtypes, given their frequency, these rare variants will only influence risk for a very small proportion of the IPF population. Moreover, our common variant analysis highlights the importance of telomerase maintenance, host defense, and cell-cell adhesion genes in the development of IPF, and our overall analysis estimated IPF heritability to be 32%. In aggregate, our results have narrowed the focus of IPF genetics to a few well-established rare variants and replicated common variants.

We found that the estimated SNP-heritability for IPF (based on all measured rare and common variants) was 32% (s.e. 3%). This estimate is similar to our previous estimates of 28% (s.e. 2%) to 31% (s.e. 3%), which were based only on common variants, excluding the *MUC5B* variant. (8) Based on these results, we hypothesize that the majority of IPF heritability can be explained by common genetic variation. However, larger sequencing datasets are needed to

explicitly estimate the contribution of rare variation to the overall heritability of IPF.

Additionally, given the relatively high heritability of IPF, common variants could be used to identify early interstitial lung disease, especially among unaffected family members.(39, 40)

Although early interstitial lung disease is known to have a poor prognosis,(40-42) screening guidelines for early interstitial lung disease have not been established and therapeutic intervention for early interstitial lung disease has not been studied.

While identifying common IPF risk variants was not a primary aim of this study, preliminary analyses that included common variants identified two previously unreported loci that were significantly associated with IPF. Within these loci, the variants with the strongest associations include an indel in *MCL1* (an apoptosis regulator in the BCL2 family at 1q21.2) and an indel in a lncRNA gene at 16p13.3. These indel variants may not have been well represented by the markers included in previous GWA studies, and will require validation in an independent cohort.

This is the first whole-genome sequencing study of IPF, with comprehensive assessment of rare variant associations outside of the exome. However, this study also has some limitations. While this study included one of the largest collections of IPF patients to date, the identification of rare variants is highly dependent on sample size, and lower frequency IPF risk alleles could possibly be identified by larger studies. Consequently, extremely rare variants in genes, such as *TERC*, *TINF2*, *ABCA3*, and the surfactant protein genes, previously identified through targeted candidate gene studies may play a role in the heritability of IPF, however, given their frequency, were not identified in our study population and will only influence risk for a very small proportion of the IPF population. In addition, the minor allele frequency

threshold used to define rare variants ($MAF \leq 1\%$) is somewhat arbitrary, and the power to identify aggregated variant sets that are associated with IPF will depend on the distribution of allele frequencies among risk variants included in a set. This study was also limited to subjects of European ancestry, and there are likely different rare variants that influence IPF risk in populations of other ancestries. Finally, the effect of rare variants may depend on age, sex, family history of disease, or other common variant genotypes. Additional analyses will be required to understand how interactions among genetic and non-genetic risk factors contribute to the etiology of IPF.

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TABLES AND FIGURES

Table 1: Summary of discovery and validation cohorts by case-control status

| | Discovery | | Validation | |
|---------------------------|-----------------------|-----------------|-----------------|-----------------|
| | Cases | Controls | Cases | Controls |
| N | 1264 | 1257 | 916 | 1200 |
| Sequencing Center | Washington University | Broad Institute | Broad Institute | Broad Institute |
| Age, Mean (SD) | 65.5 (9.4) | 59.4 (6.3) | 67.3 (9.1) | 70.0 (9.4) |
| Male, N (%) | 895 (70.8%) | 676 (53.8%) | 609 (66.5%) | 792 (66.0%) |
| Ever smoker, N (%) | 783 (68.1%) | 1198 (95.3%) | 563 (67.5%) | 742 (61.8%) |

Table 2: Significant rare variant sets identified in the discovery cohort

| Variant Filtering, Aggregation Unit | Gene/Window | P-value, discovery cohort | P-value, validation cohort [§] |
|---|---|---------------------------|---|
| Loss-of-function, by gene [*] | <i>UNC93A</i> | 3.44x10 ⁻⁴ | 0.346 |
| | <i>NFX1</i> | 4.67x10 ⁻⁴ | 0.029 |
| | <i>ALOX15B</i> | 2.81x10 ⁻⁵ | 0.705 |
| | <i>RTEL1</i> | 1.11x10 ⁻⁴ | 2.53x10 ⁻³ |
| | <i>RTEL1-TNFRSF6B</i> | 3.49x10 ⁻⁵ | 0.106 |
| Loss-of-function or missense, by gene [†] | <i>TERT</i> | 3.25x10 ⁻¹⁶ | 9.39x10 ⁻⁸ |
| | <i>RTEL1</i> | 7.49x10 ⁻⁹ | 0.021 |
| | <i>SPDL1</i> | 2.73x10 ⁻⁵ | 0.016 |
| All rare, by window [‡] | Chr11:1284193-1287310 (<i>TOLLIP</i>) | 1.45x10 ⁻⁷ | 2.89x10 ⁻⁶ |
| [*] Genes included in validation analysis based on p<5x10 ⁻⁴ [†] Genes included in validation analysis based on genome-wide significance (p<3.2x10 ⁻⁶) or p<5x10 ⁻⁵ for previously reported genes [‡] Genes included in validation analysis based on genome-wide significance (p<3.2x10 ⁻⁸) or p<5x10 ⁻⁷ for previously reported genes [§] Significance in validation cohort assessed at p<5.5x10 ⁻³ | | | |

Table 3: Meta-analysis p-values for rare variant sets included in validation analysis, with and without adjustment for common variants

| Variant Filtering, Aggregation Unit | Chr | Gene/Window | Variant Set P-value | Common Variant Covariate | Variant Set P-value, adjusted for common variant |
|---------------------------------------|---|--|------------------------|--------------------------|--|
| Loss-of-function, by gene | 6 | <i>UNC93A</i> (<i>ENSG00000112494</i>) | 2.22x10 ⁻³ | rs2076295 | 2.18x10 ⁻³ |
| | | | | rs35705950 | 9.70x10 ⁻⁴ |
| | 9 | <i>NFX1</i> (<i>ENSG00000086102</i>) | 2.84x10 ⁻⁴ | rs35705950 | 9.70x10 ⁻³ |
| | 17 | <i>ALOX15B</i> (<i>ENSG00000179593</i>) | 1.28x10 ⁻² | rs1981997 | 1.56x10 ⁻² |
| | | | | rs35705950 | 1.73x10 ⁻² |
| | 20 | <i>RTEL1</i> (<i>ENSG00000258366</i>) | 4.25x10 ⁻⁶ | rs35705950 | 1.05x10 ⁻⁵ |
| rs41308092 | | | | 4.22x10 ⁻⁶ | |
| 20 | <i>RTEL1-TNFRSF6B</i> (<i>ENSG00000026036</i>) | 9.12x10 ⁻⁵ | rs35705950 | 6.71x10 ⁻⁵ | |
| Loss-of-function or missense, by gene | 5 | <i>TERT</i> (<i>ENSG00000164362</i>) | 2.74x10 ⁻²¹ | rs4449583 | 3.47x10 ⁻²¹ |
| | | | | rs35705950 | 2.39x10 ⁻¹⁹ |
| | 5 | <i>SPDL1</i> (<i>ENSG00000040275</i>) | 4.81x10 ⁻⁶ | rs4449583 | 5.14x10 ⁻⁶ |
| | | | | rs35705950 | 1.43x10 ⁻⁶ |
| | 20 | <i>RTEL1</i> (<i>ENSG00000258366</i>) | 1.00x10 ⁻¹⁰ | rs35705950 | 3.83x10 ⁻¹¹ |
| | | | | rs41308092 | 6.11x10 ⁻¹¹ |
| All rare, by window | 11 | 11:1284193-1287310 (<i>TOLLIP</i>) | 1.04x10 ⁻¹¹ | rs35705950 | 1.00 |

Figure 1: Manhattan plot for rare loss-of-function variants (A) and loss-of-function/missense variants (B), aggregated by gene, in the discovery cohort. Horizontal lines represent the genome-wide significance thresholds adjusted for the effective number of tests (solid), and the total number of tests (dashed).

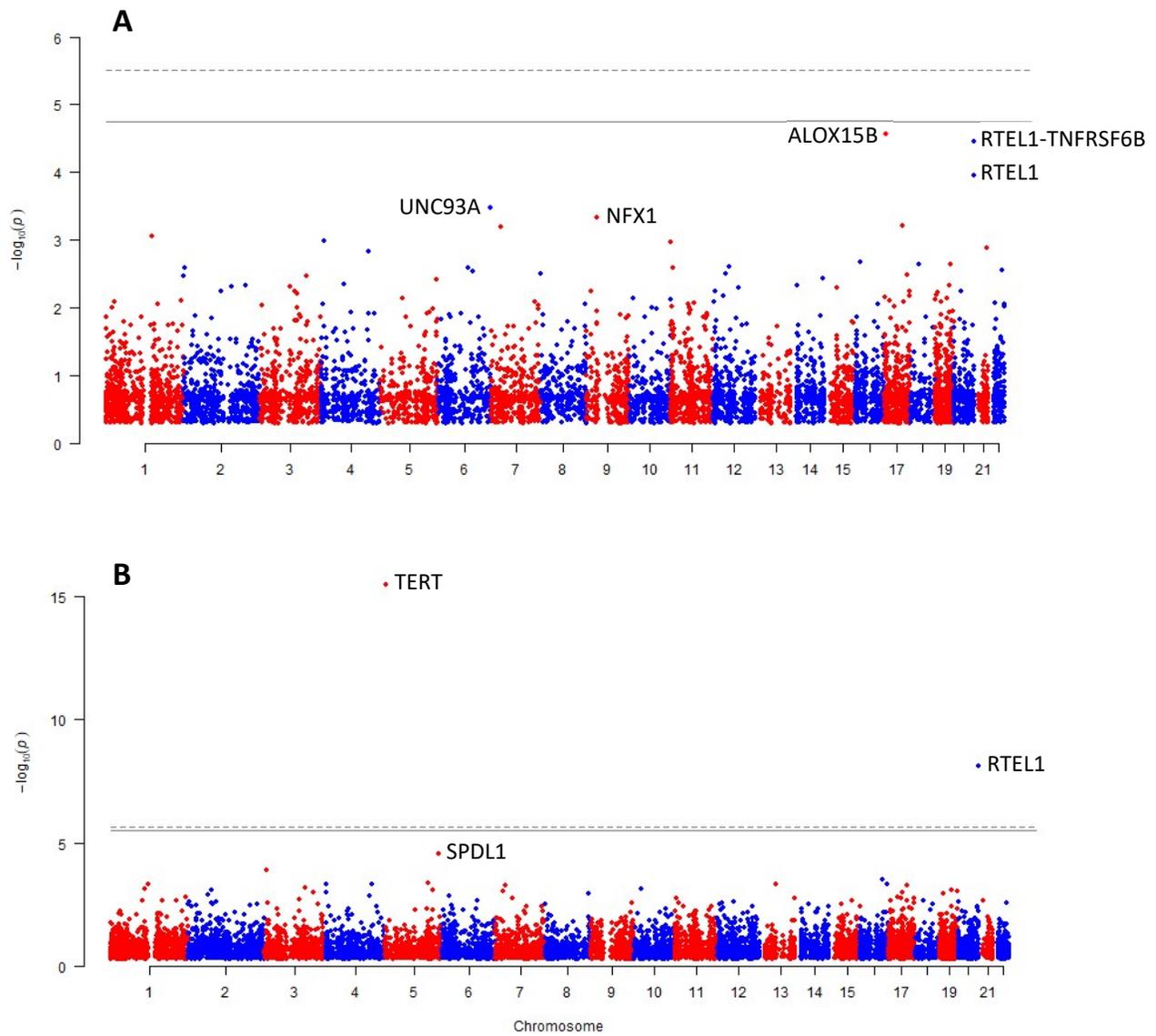


Figure 2: Manhattan plots for meta-analysis with rare loss-of-function variants (A) and loss-of-function/missense variants (B). Horizontal lines represent the genome-wide significance thresholds adjusted for the total number of tests

